## Multicollinearity

#### Class 11

### Wonmun Shin (wonmun.shin@sejong.ac.kr)

Department of Economics, Sejong University

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# Perfect Multicollinearity

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- **Perfect multicollinearity**: Existence of exact linear relationship(s) among independent variables
  - For the *K*-variable regression involving explanatory variables  $X_{1i}, X_{2i}, \dots, X_{Ki}$ (where  $X_{1i} = 1$  for all *i* to allow for the intercept term), an exact linear relationship is said to exist if the following condition is satisfied:

$$\lambda_1 X_{1i} + \lambda_2 X_{2i} + \dots + \lambda_K X_{Ki} = 0$$

where  $\lambda_1, \lambda_2, \cdots, \lambda_K$  are constants such that not all of them are zero simultaneously.

• Assume that  $\lambda_2 \neq 0$ , then

$$X_{2i} = -\frac{\lambda_1}{\lambda_2} X_{1i} - \frac{\lambda_3}{\lambda_2} X_{3i} - \dots - \frac{\lambda_K}{\lambda_2} X_{Ki}$$

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## Perfect Multicollinearity [cont'd]

- Consider the three variable model:  $Y_i = \beta_1 + \beta_2 X_{2i} + \beta_3 X_{3i} + e_i$
- Suppose  $X_{2i} = 3X_{3i}$ , which means there is exact linear relationship between  $X_{2i}$  and  $X_{3i}$ .

• Then,

$$Y_{i} = \beta_{1} + \beta_{2}X_{2i} + \beta_{3}X_{3i} + e_{i}$$
  
=  $\beta_{1} + \beta_{2}(3X_{3i}) + \beta_{3}X_{3i} + e_{i}$   
=  $\beta_{1} + (3\beta_{2} + \beta_{3})X_{3i} + e_{i}$ 

• Now, we cannot estimate either  $\beta_2$  and  $\beta_3$  separately!!

- We can estimate only the linear combination of two coefficients, *i.e.*  $3\hat{\beta}_2 + \hat{\beta}_3$ .
- There is no unique value for  $\hat{\beta}_2$  (or  $\hat{\beta}_3$ ).
- This problem is usually called as Identification Problem.

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## Perfect Multicollinearity [cont'd]

• Recall the OLS estimator in the three-variable regression model:

$$\hat{\beta}_{2} = \frac{\sum x_{2i} y_{i} \sum x_{3i}^{2} - \sum x_{3i} y_{i} \sum x_{2i} x_{3i}}{\sum x_{2i}^{2} \sum x_{3i}^{2} - (\sum x_{2i} x_{3i})^{2}}$$

• Suppose that  $X_{2i} = 3X_{3i} \implies$  indeterminate expression

$$\hat{\beta}_{2} = \frac{3\sum x_{3i}y_{i}\sum x_{3i}^{2} - 3\sum x_{3i}y_{i}\sum x_{3i}^{2}}{3^{2}\sum x_{3i}^{2}\sum x_{3i}^{2} - (3\sum x_{3i}^{2})^{2}} = \frac{0}{0}$$

- Why do we obtain this result?
  - Meaning of  $\beta_2$ : The rate of change in the average value of Y as  $X_2$  changes by a unit, holding  $X_3$  constant
  - But if X<sub>2</sub> and X<sub>3</sub> are perfectly collinear, there is no way X<sub>3</sub> can be kept constant.
  - As X<sub>2</sub> changes, so does X<sub>3</sub> by 3: There is no way of disentangling the separate influences of X<sub>2</sub> and X<sub>3</sub> from the given sample.

# Near Multicollinearity

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## Near Multicollinearity

- Originally, multicollinearity meant the existence of a perfect linear relationship among regressors.
- Today, however, the term multicollinearity is used in a broader sense.
  - Includes the case of perfect multicollinearity as well as the case where the X variables are intercorrelated but not perfectly.

$$\lambda_1 X_{1i} + \lambda_2 X_{2i} + \dots + \lambda_K X_{Ki} + v_i = 0$$

where  $v_i$  is a stochastic error term.

• Assume that  $\lambda_2 \neq 0$ , then

$$X_{2i} = -\frac{\lambda_1}{\lambda_2} X_{1i} - \frac{\lambda_3}{\lambda_2} X_{3i} - \dots - \frac{\lambda_K}{\lambda_2} X_{Ki} - \frac{1}{\lambda_2} v_i$$

which shows that  $X_2$  is not an exact linear combination of other X's because it is also determined by the stochastic error term  $v_i$ .

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## Near Multicollinearity [cont'd]

X <sub>2i</sub>	X <sub>3i</sub>	X*3i
10	50	52
15	75	75
18	90	97
24	120	129
30	150	152

•  $X_{3i} = 5X_{2i}$ : There is perfect multicollinearity between  $X_2$  and  $X_3$ . ( $r_{23} = 1$ )

- $X_{3i}^*$  was created from  $X_3$  by simply adding the random numbers to it .
  - Now there is no longer perfect multicollinearity between  $X_{2i}$  and  $X_{3i}^*$ .
  - However, the two variables are highly correlated!  $(r_{23^*} = 0.9959)$
  - In this case, there is Near (Perfect) Multicollinearity (or High but Imperfect Multicollinearity).

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• Practical example:

$$Y_i = \beta_1 + \beta_2 X_{2i} + \beta_3 X_{3i} + e_i$$

- Y<sub>i</sub>: consumption expenditure
- X<sub>2i</sub>: income
- X<sub>3i</sub>: wealth
- $X_{3i}$  may have close positive relationship with  $X_{2i}!$

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# Consequences of Multicollinearity

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1. Although BLUE, the OLS estimators may have *large variances*.

- (Near) multicollinearity does not violate the classical assumptions.
- Gauss-Markov Theorem  $\Rightarrow$  OLS estimators are BLUE!
- However, they have large variances, making precise estimation difficult.
- For the three-variable regression model, we can obtain

$$Var\left(\hat{eta}_{2}
ight)=rac{\sigma^{2}}{\sum x_{2i}^{2}\left(1-r_{23}^{2}
ight)}$$

- As  $r_{23}$  tends toward 1 or -1 (*i.e.* as collinearity increases), the variance of  $\hat{\beta}_2$  increases.
- Extremely, when  $r_{23} = 1$  (or -1),  $Var(\hat{\beta}_2)$  is infinite ( $\rightarrow$  identification problem).

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#### 2. Insignificant t ratios

- Because of large variance (**Consequence 1**), the *t* ratio of one or more coefficients tends to be statistically insignificant.
- Recall, to test  $H_0: \beta_2 = 0$ , we use the *t* ratio.

$$t = \frac{\hat{\beta}_2}{s.e.\left(\hat{\beta}_2\right)}$$

- As we have seen, in cases of high collinearity, the estimated standard errors increase dramatically.
- Therefore, t value becomes smaller  $\longrightarrow$  One will increasingly accept the null hypothesis!
- Probability of Type II error increases (or low power of test).
  - cf. Type II error: Error of not rejecting  $H_0$  when  $H_0$  is false.
- 3. Wider confidence interval

## Consequences of Multicollinearity [cont'd]

- 4. Significant *F* ratio but few significant *t* ratios
  - Although the t ratio of one or more coefficients is statistically insignificant,  $\frac{R^2}{C}$  (overall measure of goodness of fit) can be very high.
  - Consider the K-variable linear regression model:

$$Y_i = \beta_1 + \beta_2 X_{2i} + \dots + \beta_K X_{Ki} + e_i$$

• *F*-test of overall significance of coefficients:  $H_0: \beta_2 = \cdots = \beta_K = 0$ 

$$F = \frac{R^2}{1 - R^2} \cdot \frac{n - K}{K - 1} \sim F(K - 1, n - K)$$

- F can be very large even if individual t ratio is insignificant!
- Under multicollinearity, explanatory variables can be jointly significant even if each of them is individually insignificant.
  - Remember the distinction between a joint test and an individual test.
- 5. OLS estimators and their standard errors may be **sensitive** to small changes in the data.

## Detection and Remedies of Multicollinearity

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- How do we know that multicollinearity is present in any given situation?
  - Multicollinearity is essentially a sample phenomenon (arising out of non-experimental data collected in most social sciences).
  - Therefore, we do not have one unique method of detecting it or measuring its strength.
- 1. High  $R^2$  (or significant F ratio) but few significant *t*-ratios.
  - If  $R^2$  is high (usually, in excess of 0.8) and individual t test show that none or very few coefficients are significant, then there may be multicollinearity problem.
  - **Problem:** Although this diagnostic is sensible, its disadvantage is that it is too strong.
    - Note that multicollinearity is considered as "harmful" only when all of the influences of regressors on Y cannot be disentangled.

- 2. High pairwise correlation(s) among explanatory variables
  - If the pairwise correlation coefficient between two regressors is high (in excess of 0.8, in absolute value), then multicollinearity is a serious problem.
  - **Problem:** Multicollinearity can exist even though the pairwise correlations are comparatively low (less than 0.5, in absolute value).
    - High pairwise correlations are a sufficient condition but not a necessary condition for the existence of multicollinearity.
- 3. Auxiliary regressions
  - One way of finding out which X variable is related to other X variables is to regress each X<sub>ki</sub> on the remaining X variables.
  - Each one of these regression is called an **auxiliary regression** (auxiliary to the main regression of Y on X's).

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### 1. Do nothing

- Multicollinearity is essentially a data problem, and sometimes we have no choice over data available.
- Also, it is not that all coefficients in a regression model are insignificant.
  - Moreover, significant F ratio means the model is overall significant.
  - Depending on the objective of analysis, multicollinearity is not that problematic.

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### 2. A Priori information (about parameters)

• **Example 1:** Suppose we know that  $\beta_3 = 0.1\beta_2$  (given a priori information).

$$Y_{i} = \beta_{1} + \beta_{2}X_{2i} + \beta_{3}X_{3i} + e_{i}$$
  
=  $\beta_{1} + \beta_{2}(X_{2i} + 0.1X_{3i}) + e_{i}$ 

- Once we obtain  $\hat{\beta}_2,$  we can estimate  $\hat{\beta}_3$  from the relationship between  $\beta_2$  and  $\beta_3.$
- Example 2: Cobb-Douglas production function

$$Y_i = \beta_1 L_i^{\beta_2} K_i^{\beta_3} e^{e_i}$$
$$\implies \ln Y_i = \ln \beta_1 + \beta_2 \ln L_i + \beta_3 \ln K_i + e_i$$

• Suppose we know CRTS:  $\beta_2+\beta_3=1$ 

$$\ln Y_i = \ln \beta_1 + \beta_2 \ln L_i + (1 - \beta_2) \ln K_i + e_i$$
$$\implies \ln (Y_i/K_i) = \ln \beta_1 + \beta_2 \ln (L_i/K_i) + e_i$$

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### 3. Dropping a variable(s)

- The simplest remedy: dropping one of the collinear variables
- But, we may be committing a specification error of omitting relevant variables (→ *biased* and *inconsistent* estimator)

### 4. Transformation of variables: First difference

- One reason for high multicollinearity : Variables tent to move in the same direction over time (eg. income and wealth)
- Although the variables may be highly correlated, there is no a priori reason to believe their first differences will also be highly correlated.
- But, it might introduce autocorrelation of errors.
- Example:

$$\underbrace{Y_t - Y_{t-1}}_{\triangle Y_t} = \beta_2 \underbrace{(X_{2t} - X_{2t-1})}_{\triangle X_{2t}} + \beta_3 \underbrace{(X_{3t} - X_{3t-1})}_{\triangle X_{3t}} + \underbrace{(e_t - e_{t-1})}_{\nu_t}$$

- 5. Transformation of variables: Ratio transformation
  - Example : Y is consumption,  $X_2$  is GDP,  $X_3$  is population
    - GDP and population grow over time.  $\rightarrow$  They are likely to be correlated.
    - One solution is to express the model on a per capita basis:

$$\frac{Y_t}{X_{3t}} = \beta_1 \left(\frac{1}{X_{3t}}\right) + \beta_2 \left(\frac{X_{2t}}{X_{3t}}\right) + \beta_3 + \left(\frac{e_t}{X_{3t}}\right)$$

• But, it might introduce heteroskedastic errors.

#### 6. Additional or new data

- Again, multicollinearity is a sample feature.
- Simply increasing the size of the sample may attenuate the collinearity problem.

$$n \Uparrow \rightarrow Var\left(\hat{eta}_2
ight) = rac{\sigma^2}{\sum x_{2i}^2 \left(1 - r_{23}^2
ight)} \Downarrow$$